

Second Order Properties of Distribution Tails and Estimation of Tail Exponents in Random Difference Equations

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Random Difference Equations (RDEs)

In one dimension, RDE is given by

$$X_n = A_n X_{n-1} + B_n, \quad X_0 = x, \quad n \geq 1,$$

where (A_n, B_n) are typically assumed to be i.i.d. vectors. Its stationary solution, also called RDE, satisfies

$$X \stackrel{d}{=} AX + B$$

with $(A, B) \stackrel{d}{=} (A_1, B_1)$ independent of X . Focus on $A, B, X > 0$.

Theorem (Kesten (1973))

Under mild assumptions on A and B , it is known that

$$P(X > x) \sim cx^{-\alpha},$$

for $\alpha > 0$ satisfying $EA^\alpha = 1$.

Examples of RDEs

- Beta prime distribution $\beta(a, b)$: The only known RDEs X with explicit power-law tail (tail exponent is b),

$$\beta(a, b) \stackrel{d}{=} \frac{1}{\text{Beta}(b, a)} - 1.$$

- (Squares of) ARCH(1) model in Finance

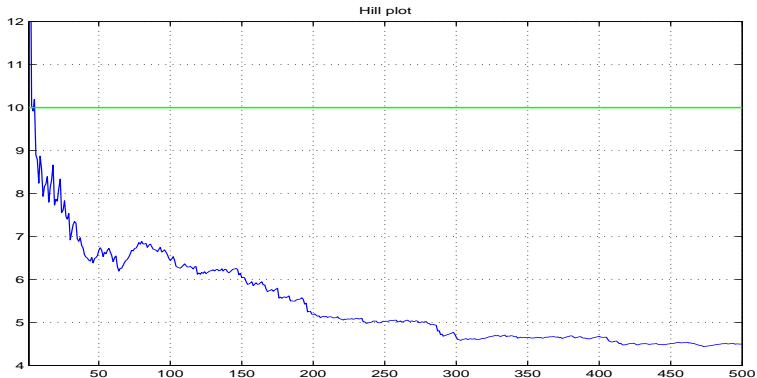
$$\xi_n^2 = \lambda \epsilon_n^2 \xi_{n-1}^2 + \beta \epsilon_n^2.$$

- MC (Multiplicative Cascades) in Physics with larger tail exponent is expected in practice.

Statement of problem

Huge bias from Hill plots for larger tail exponent.

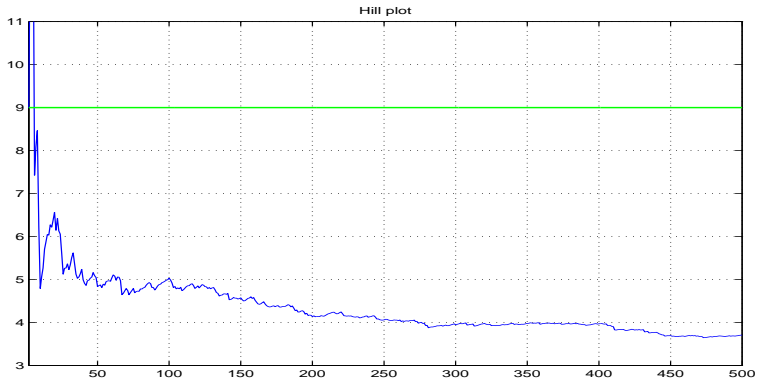
Hill plot for 5000 **independent** realizations of ARCH(1) model with $\alpha = 10$.



Statement of problem

Huge bias from Hill plots for larger tail exponent.

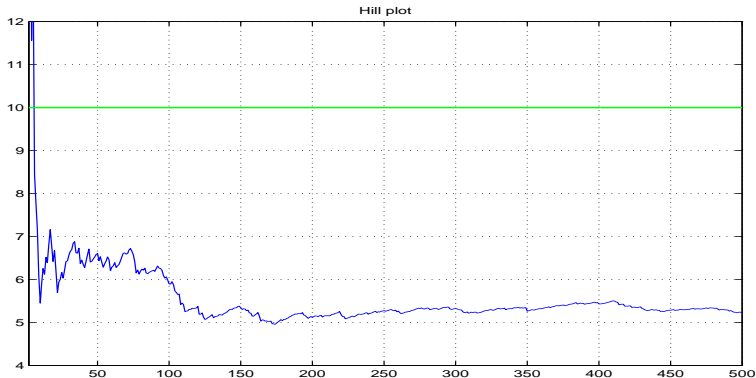
Hill plot for 5000 **independent** $\beta(5, 9)$ distribution with $\alpha = 9$.



Statement of problem

Huge bias from Hill plots for larger tail exponent.

Hill plot for 5000 **independent** realizations of MC with $\alpha = 10$.



Possible explanations

Because dependence is removed two explanations are possible:

- X is replaced by X_N in practice and the approximation is poor. This, however, is **not the case**.
- The distribution tail of X deviates from exact power-law tail.

Second order regular variation (2RV)

2RV is a natural framework to examine bias.

The distribution tail $P(X > x)$ is 2RV with first order parameter $\alpha > 0$ and second order parameter $\rho < 0$, if for suitable $g(x)$ and any $a > 0$,

$$\lim_{x \rightarrow \infty} \frac{x^\alpha P(X > x) - (ax)^\alpha P(X > ax)}{g(x)} = \frac{a^\rho - 1}{\rho}.$$

In practice, this is replaced by

$$P(X > x) = c_1 x^{-\alpha} + c_2 x^{-\alpha+\rho} + o(x^{-\alpha+\rho}).$$

2RV-like properties for RDEs

2RV for RDEs appears an open and difficult problem.
Nevertheless, one can show:

Theorem 1 (weaker form of 2RV for RDEs)

Under mild assumptions on A , B and X , one has

$$\int_x^\infty (P(X > u) - P(AX > u)) du \sim Cx^{-\alpha} \quad (\star)$$

Remark. This result suggests that

$$P(X > x) - P(AX > x) \sim C\alpha x^{-\alpha-1}$$

or

$$\lim_{x \rightarrow \infty} \frac{x^\alpha P(X > x) - x^\alpha (EA^\alpha)^{-1} P(X > A^{-1}x)}{x^{-1}} = C\alpha,$$

which can be viewed as 2RV at random $a = 1/A$.

2RV-like properties for RDEs

Remark. Theorem 1 is proved just by using Kesten's result and the form of RDEs.

On a more practical side, one has:

Theorem 2 (important relation to practical 2RV)

Under mild assumptions on X , A and B , the result (★) in Theorem 1 is consistent with

$$P(X > x) = c_1 x^{-\alpha} + c_2 x^{-\alpha+\rho} + o(x^{-\alpha+\rho}),$$

only when

$$\rho = -1.$$

Extensions and open questions

Extensions:

- RDEs with real-valued coefficients.
- Multidimensional RDEs with nonnegative matrix coefficients.
Examples include GARCH model.

Open questions:

- Exact 2RV.
- Relevance in practice.

Estimation with known ρ

The results above suggest that for RDE X

$$P(X > x) = c_1 x^{-\alpha} + c_2 x^{-\alpha+\rho} + o(x^{-\alpha+\rho})$$

with $\rho = -1$.

- Large literature on estimation when ρ is *unknown*.
Feuerverger and Hall (1999), Beirlant, Dierckx, Goegebeur and Matthys (1999), Beirlant, Dierckx, Guillou and Stărică (2002), Gomes and Martins (2002) to name a few.
- For known ρ , what would be the best estimator?

Number of possible estimators

RK2 Rank-based least squares estimator based on

$$\log P(X > x) \approx \log c_1 - \alpha \log x + c_2/c_1 x^{-1}$$

Replacing x by order statistic $X_{(n-i+1)}$ gives

$$\log(i/n) \approx \log c_1 - \alpha \log X_{(n-i+1)} + c_2/c_1 X_{(n-i+1)}^{-1}.$$

QQ2 Reversing the role of $\log P(X > x)$ and $\log x$.

JK, P Generalized jackknife estimators adapting Gomes, Martins and Neves (2000) and Peng (1998).

FH Feuerverger and Hall (1999) estimator with known ρ .

ML Maximum likelihood estimator based on the exact tail
 $P(X > x) = c_1 x^{-\alpha} + c_2 x^{-\alpha-1}.$

Properties of proposed estimators

Theorem 3 (asymptotic normality)

Under suitable assumptions,

$$\sqrt{k}(\hat{\alpha} - \alpha) \xrightarrow{d} \mathcal{N}(0, \sigma^2),$$

where k is the number of upper order statistics used and asymptotic variance

$$\sigma_{RK2}^2 = \sigma_{QQ2}^2 = 2\alpha^2(\alpha + 1)^2,$$

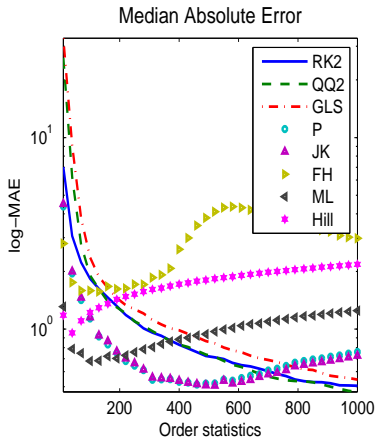
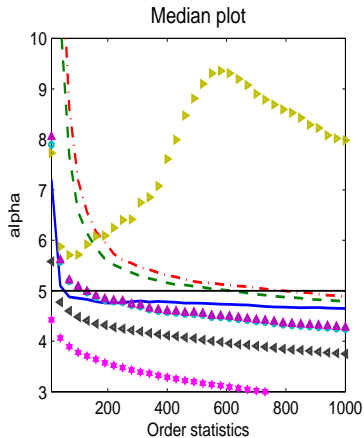
$$\sigma_{JK}^2 = \sigma_P^2 = \alpha^2((\alpha + 1)^2 + \alpha^2),$$

$$\sigma_{FH}^2 = \sigma_{ML}^2 = \alpha^2(\alpha + 1)^2.$$

In simulations: proposed estimators successfully provide bias correction. RK2 estimator appears to perform best in practice.

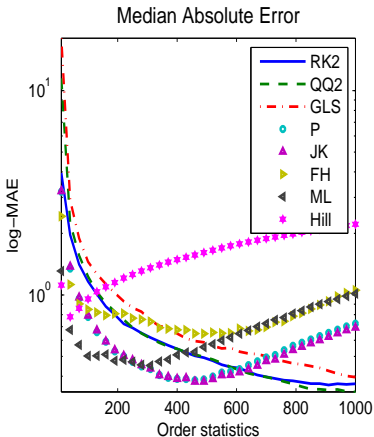
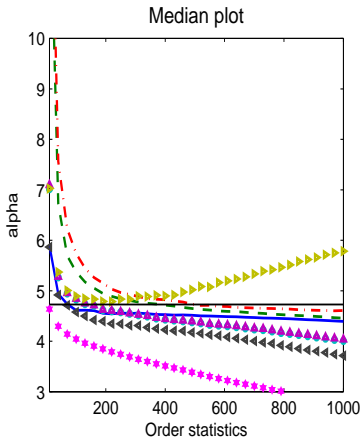
Comparison in simulations

Beta prime distribution with $\alpha = 5$, $T = 5,000$.



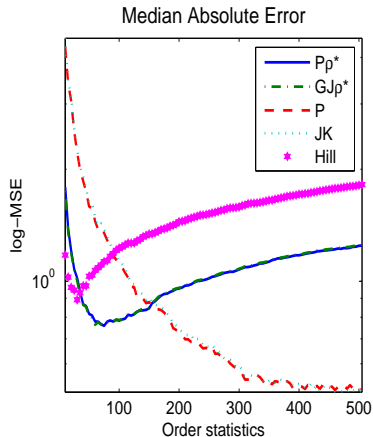
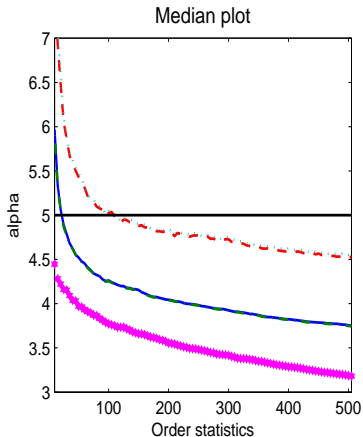
Comparison in simulations

ARCH(1) model with $\alpha = 4.73$, $T = 5,000$.



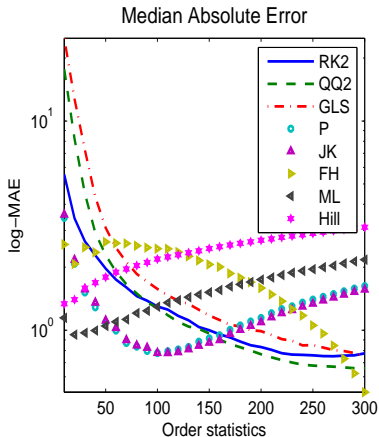
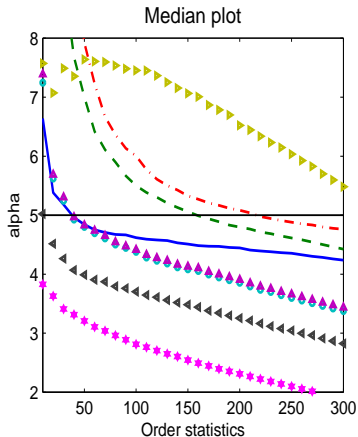
Comparison in simulations

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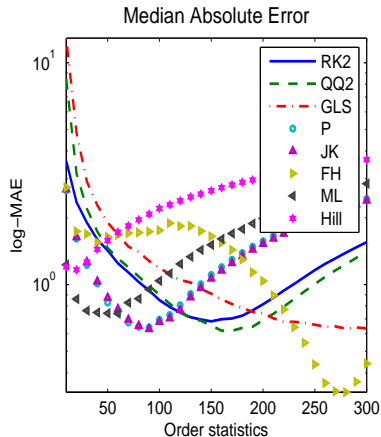
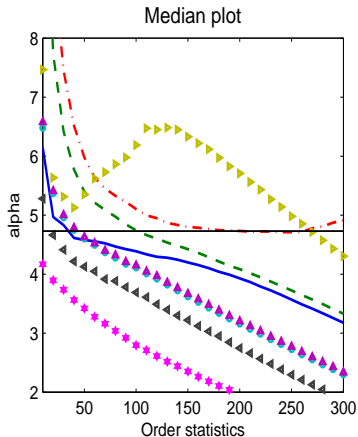
Comparison in simulations

Beta prime distribution with $\alpha = 5$, $T = 500$.



Comparison in simulations

ARCH(1) model with $\alpha = 4.73$, $T = 500$.



References

- 1 Baek, C., Pipliras, V., Wendt, H. & Abry, P. (2009), 'Second order properties of distribution tails and estimation of tail exponents in random difference equations', To appear in Extremes.
- 2 Baek, C. & Pipliras, V. (2009), 'Estimation of parameters in heavy-tailed distribution when its second order tail parameter is known', Preprint.

Thank you!