

6.4.6 a) In order for  $\alpha$  to be 0.07,  $P(60 - \bar{y}^* \leq \bar{Y} \leq 60 + \bar{y}^* \mid \mu = 60) = 0.07$ . Equivalently,

$$P\left(\frac{60 - \bar{y}^* - 60}{8.0/\sqrt{36}} \leq \frac{\bar{Y} - 60}{8.0/\sqrt{36}} \leq \frac{60 + \bar{y}^* - 60}{8.0/\sqrt{36}}\right) = P(-0.75\bar{y}^* \leq Z \leq 0.75\bar{y}^*) = 0.07. \text{ But}$$

$$P(-0.09 \leq Z \leq 0.09) = 0.07, \text{ so } 0.75\bar{y}^* = 0.09, \text{ which implies that } \bar{y}^* = 0.12.$$

b)  $1 - \beta = P(\text{reject } H_0 \mid H_1 \text{ is true}) = P(59.88 \leq \bar{Y} \leq 60.12 \mid \mu = 62) =$

$$P\left(\frac{59.88 - 62}{8.0/\sqrt{36}} \leq Z \leq \frac{60.12 - 62}{8.0/\sqrt{36}}\right) = P(-1.59 \leq Z \leq -1.41) = 0.0793 - 0.0559 = 0.0234.$$

c) For  $\alpha = 0.07$ ,  $\pm z_{\alpha/2} = \pm 1.81$  and  $H_0$  should be rejected if  $\bar{y}$  is either

$$1) \leq 60 - 1.81 \cdot \frac{8.0}{\sqrt{36}} = 57.50 \text{ or } 2) \geq 60 + 1.81 \cdot \frac{8.0}{\sqrt{36}} = 62.41. \text{ Suppose } \mu = 62. \text{ Then}$$

$$1 - \beta = P(\bar{Y} \leq 57.59 \mid \mu = 62) + P(\bar{Y} \geq 62.41 \mid \mu = 62) = P(Z \leq -3.31) + P(Z \geq 0.31) = 0.0005 + 0.3783 = 0.3788.$$

6.4.8 If  $n = 45$ ,  $H_0$  will be rejected when  $\bar{y}$  is either 1)  $\leq 10 - 1.96 \cdot \frac{4}{\sqrt{45}} = 8.83$  or 2)  $\geq 10 +$

$$1.96 \cdot \frac{4}{\sqrt{45}} = 11.17. \text{ When } \mu = 12, \beta = P(\text{accept } H_0 \mid H_1 \text{ is true}) = P(8.83 \leq \bar{Y} \leq 11.17 \mid \mu =$$

$$12) = P\left(\frac{8.83 - 12}{4/\sqrt{45}} \leq Z \leq \frac{11.17 - 12}{4/\sqrt{45}}\right) = P(-5.32 \leq Z \leq -1.39) = 0.0823. \text{ It follows that a}$$

sample of size  $n = 45$  is sufficient to keep  $\beta$  smaller than 0.20 when  $\mu = 12$ .

6.4.10 a)  $P(\text{Type I error}) = P(\text{reject } H_0 \mid H_0 \text{ is true}) = P(Y \geq 3.20 \mid \lambda = 1) =$

$$\int_{3.20}^{\infty} e^{-y} dy = 0.04.$$

b)  $P(\text{Type II error}) = P(\text{accept } H_0 \mid H_1 \text{ is true}) =$

$$P\left(Y < 3.20 \mid \lambda = \frac{4}{3}\right) = \int_0^{3.20} \frac{3}{4} e^{-3y/4} dy = \int_0^{2.4} e^{-u} du = 0.91.$$

6.4.16 If  $H_0$  is true,  $X = X_1 + X_2$  has a binomial distribution with  $n = 6$  and  $p = \frac{1}{2}$ . Therefore,

$$\alpha = P(\text{reject } H_0 \mid H_0 \text{ is true}) = P\left(X \geq 5 \mid p = \frac{1}{2}\right) = \sum_{k=5}^6 \binom{6}{k} \left(\frac{1}{2}\right)^k \left(1 - \frac{1}{2}\right)^{6-k} = 7/2^6 = 0.11.$$

$$6.4.18 \text{ a) } \alpha = P(\text{reject } H_0 \mid H_0 \text{ is true}) = P(X \leq 2 \mid \lambda = 6) = \sum_{k=0}^2 \frac{e^{-6} 6^k}{k!} = 0.062.$$

$$\text{b) } \beta = P(\text{accept } H_0 \mid H_1 \text{ is true}) = P(X \geq 3 \mid \lambda = 4) = 1 - P(X \leq 2 \mid \lambda = 4) = 1 - \sum_{k=0}^2 \frac{e^{-4} 4^k}{k!} = 1 - 0.238 = 0.762.$$

$$6.4.20 \quad \beta = P(\text{accept } H_0 \mid H_1 \text{ is true}) = P(Y < \ln 10 \mid \lambda) = \int_0^{\ln 10} \lambda e^{-\lambda y} dy = 1 - e^{-\lambda \ln 10} = 1 - 10^{-\lambda}.$$

6.4.21  $\alpha = P(\text{reject } H_0 \mid H_0 \text{ is true}) = P(Y_1 + Y_2 \leq k \mid \theta = 2)$ . When  $H_0$  is true,  $Y_1$  and  $Y_2$  are uniformly distributed over the square defined by  $0 \leq Y_1 \leq 2$  and  $0 \leq Y_2 \leq 2$ , so the joint pdf of  $Y_1$  and  $Y_2$  is a plane parallel to the  $Y_1 Y_2$ -axis at height  $\frac{1}{4}$  ( $= f_{Y_1}(y_1) \cdot f_{Y_2}(y_2) = \frac{1}{2} \cdot \frac{1}{2}$ ). By geometry,  $\alpha$  is the volume of the triangular wedge in the lower left-hand corner of the square over which  $Y_1$  and  $Y_2$  are defined. The hypotenuse of the triangle in the  $Y_1 Y_2$ -plane has the equation  $y_1 + y_2 = k$ . Therefore,  $\alpha = \text{area of triangle} \times \text{height of wedge} = \frac{1}{2} \cdot k \cdot k \cdot \frac{1}{4} = k^2/8$ . For  $\alpha$  to be 0.05,  $k = \sqrt{0.04} = 0.63$ .

6.4.22  $\alpha = P(\text{reject } H_0 \mid H_0 \text{ is true}) = P(Y_1 Y_2 \leq k^* \mid \theta = 2)$ . If  $\theta = 2$ , the joint pdf of  $Y_1$  and  $Y_2$  is the horizontal plane  $f_{Y_1, Y_2}(y_1, y_2) = \frac{1}{4}$ ,  $0 \leq y_1 \leq 2$ ,  $0 \leq y_2 \leq 2$ . Therefore,  $\alpha = P(Y_1 Y_2 \leq k^* \mid \theta = 2) = 2 \cdot \frac{k^*}{2} \cdot \frac{1}{4} + \int_{k^*/2}^2 \int_0^{k^*/y_1} \frac{1}{4} dy_2 dy_1 = \frac{k^*}{4} + \int_{k^*/2}^2 \frac{k^*}{4 y_1} dy_1 = \frac{k^*}{4} + \left( \frac{k^*}{4} \ln y_1 \Big|_{k^*/2}^2 \right) = \frac{k^*}{4} + \frac{k^*}{4} \ln 2 - \frac{k^*}{4} \ln \frac{k^*}{2}$ . By trial and error,  $k^* = 0.087$  makes  $\alpha = 0.05$ .

Note: The  $k^*$  value in 6.4.22 is incorrect. The correct value is approximately 0.0349.